Representing Uncertainty in ML

DL2 – 2024

Wilker Aziz w.aziz@uva.nl



UNIVERSITY OF AMSTERDAM Institute for Logic, Language and Computation

Hello

I am an assistant professor at the Institute for Logic, Language and Computation (University of Amsterdam). You can check some of my work here https://probabll.github.io

Stuff I typically work on include

• machine learning

approximate (Bayesian) inference, gradient estimation, normalising flows, latent variables models (e.g., VAEs),

• natural language processing

translation, text classification, question answering, transparent and interpretable models

I teach advanced topics in DL (e.g., deep generative models, approximate inference) and NLP.

Some of my classes (in collaboration with UvA colleagues) can be found at https://uvadl2c.github.io and https://probabll.github.io/ teaching/.

- Representing uncertainty in ML uncertainty — probabilistic models — tools for prescribing distributions
- Generative models (exact) autoregressive models — normalising flows
- Generative models (approximate) energy-based models — score-matching and diffusion
- Latent variable models
 exact inference variational inference

How to use this: during the class, don't go on reading everything you see on this side. I will walk you through what's needed. I hope these notes will help you when you refer back to the content in your own time.

Outline

Uncertainty

2 Probabilistic Models

Modelling Random Experiments

4 Modelling Observed Random Variables

- 5 Tools for prescribing distributions
 - Univariate
 - Multivariate

 the quality of being indeterminate as to magnitude or value; the amount of variation in a numerical result that is consistent with observation; 1. Too specific. We can be uncertain about non-numerical things.

- the quality of being indeterminate as to magnitude or value; the amount of variation in a numerical result that is consistent with observation;
- It he state of not being definitely known or perfectly clear;

- 1. Too specific. We can be uncertain about non-numerical things.
- Interesting: uncertainty as a property of the other (what we observe, contemplate or interact with). But you and I may be differently uncertain about the same thing (e.g., I don't know much about finance, you might).

- the quality of being indeterminate as to magnitude or value; the amount of variation in a numerical result that is consistent with observation;
- the state of not being definitely known or perfectly clear;
- It the state or character of being uncertain in mind; a state of doubt

- 1. Too specific. We can be uncertain about non-numerical things.
- Interesting: uncertainty as a property of the other (what we observe, contemplate or interact with). But you and I may be differently uncertain about the same thing (e.g., I don't know much about finance, you might).
- 3. My favourite: uncertainty as a property of the self (i.e., I am uncertain about stuff, so are you, on occasion we might agree but, generally, my uncertainty about stuff owes little to yours).

- the quality of being indeterminate as to magnitude or value; the amount of variation in a numerical result that is consistent with observation;
- the state of not being definitely known or perfectly clear;
- **(3)** the state or character of being uncertain in mind; a state of doubt

Two views

Event-centric uncertainty concerns stochasticity inherent to that which we contemplate or interact with.

Agent-centric uncertainty concerns our own state of knowledge about what we contemplate or interact with (independently of those being stochastic themselves).

The agent-centric view can be argued to generalise the event-centric view (\approx agents have access to the exact same information and agree to use it in the same way). (De Finetti, 1974; Lindley, 2013)

You and I observe a sequence of coin flips, we each represent our uncertainty about the next flip.

- Event-centric approach says our representations must be identical, else one or both of us is being irrational.
- Agent-centric approach says our representations need not be identical (e.g., we might possess different information about the flips and the physics of coins).

You and I read the first half of the lord of the rings, our uncertainty about the finale need not be the same.

- Event-centric: struggles with the fact that the finale isn't stochastic (we cannot have the author rewrite it).
- Agent-centric: one can always express their own uncertainty over something they lack information about.

Formal account

A representation of the state of knowledge of an agent. Most theories are built on two key frameworks:

- Possible worlds (Hintikka, 1957, 1961; Menzel, 2023): a set algebra used to represent knowledge and possibility.
- Plausibility measures (Friedman and Halpern, 1996): a tool to order propositions as to express a preference for those we are less uncertain about.

Book recommendation

• Overview of formal frameworks: Halpern (2017)

Overview of formal frameworks: Halpern (2017).

| Uncertainty | |
|-------------|--|
| | |

• You contemplate or interact with the world, but you do not know the actual state of the world.

• For example, if you are contemplating the result of rolling a six-sided die, you may represent a world (an outcome of the experiment) as a symbol w_k where k is the number of pips the die shows.

- You contemplate or interact with the world, but you do not know the actual state of the world.
- You represent a world as a symbol (or collection of attributes) and you assume the actual world ω is one of a set Ω of possible worlds.

- For example, if you are contemplating the result of rolling a six-sided die, you may represent a world (an outcome of the experiment) as a symbol w_k where k is the number of pips the die shows.
- Assuming dice always land on exactly one of their 6 faces (never on an edge), you take Ω = {w₁, w₂, w₃, w₄, w₅, w₆} to represent all possible worlds.

- You contemplate or interact with the world, but you do not know the actual state of the world.
- You represent a world as a symbol (or collection of attributes) and you assume the actual world ω is one of a set Ω of possible worlds.
- You use subsets of Ω to convey knowledge and possibility (or propositions).

- For example, if you are contemplating the result of rolling a six-sided die, you may represent a world (an outcome of the experiment) as a symbol w_k where k is the number of pips the die shows.
- Assuming dice always land on exactly one of their 6 faces (never on an edge), you take Ω = {w₁, w₂, w₃, w₄, w₅, w₆} to represent all possible worlds.
- If you speculate the actual world ω is an odd number you make the proposition that ω is in A_{odd} = {w ∈ Ω : odd(w)} = {w₁, w₃, w₅}.

- You contemplate or interact with the world, but you do not know the actual state of the world.
- You represent a world as a symbol (or collection of attributes) and you assume the actual world ω is one of a set Ω of possible worlds.
- You use subsets of Ω to convey knowledge and possibility (or propositions).
- The set Σ of subsets of Ω you acknowledge as possible form a σ-algebra (i.e., closed under complementation and countable union/intersection).

- For example, if you are contemplating the result of rolling a six-sided die, you may represent a world (an outcome of the experiment) as a symbol w_k where k is the number of pips the die shows.
- Assuming dice always land on exactly one of their 6 faces (never on an edge), you take Ω = {w₁, w₂, w₃, w₄, w₅, w₆} to represent all possible worlds.
- If you speculate the actual world ω is an odd number you make the proposition that ω is in A_{odd} = {w ∈ Ω : odd(w)} = {w₁, w₃, w₅}.
- If you know the actual world ω is odd or prime and not one that can be decomposed as a sum of two other distinct worlds, then you know ω is in: $(\{1,3,5\} \cup \{2,3,5\}) \cap (\Omega \setminus \{3,5\}) = \{1,2\}$

- You contemplate or interact with the world, but you do not know the actual state of the world.
- You represent a world as a symbol (or collection of attributes) and you assume the actual world ω is one of a set Ω of possible worlds.
- You use subsets of Ω to convey knowledge and possibility (or propositions).
- The set Σ of subsets of Ω you acknowledge as possible form a σ-algebra (i.e., closed under complementation and countable union/intersection).

Possible worlds – examples

A2 resolves named entities to their Wikidata entry:

- a world is a symbol of the kind *w_e*, *e* is a unique id (e.g., Q76 is Barack Obama);
- the universe is $\Omega = \{w_e : e \in Wikidata\}$
- {w_e : citizenship(e, USA) ∧ politician(e)} ⊆ Ω is the proposition that the actual world is a USA politician

Possible worlds – examples

A2 resolves named entities to their Wikidata entry:

- a world is a symbol of the kind *w_e*, *e* is a unique id (e.g., Q76 is Barack Obama);
- the universe is $\Omega = \{w_e : e \in Wikidata\}$
- {w_e : citizenship(e, USA) ∧ politician(e)} ⊆ Ω is the proposition that the actual world is a USA politician

A3 responds to a question on a chat-like tool

- a world is a symbol of the kind w_s, s is a string of symbols from some vocabulary Σ;
- \bullet the universe is $\Omega=\Sigma^*$
- {w_s : presidentof(JoeBiden, USA) = parse(s)} ⊆ Ω is the proposition that the actual world is a sentence that expresses the relation presidentof(JoeBiden, USA).

8 / 66

• Your uncertainty is qualitatively different for different propositions.

You are at least as uncertain about A_{odd,prime} = A_{odd} ∩ A_{prime} as you are about A_{odd}, simply because the later includes the former.

- Your uncertainty is qualitatively different for different propositions.
- You attach an uncertainty qualifier to each proposition in Σ . A qualifier is anything you can order (e.g., fractions of a stick, numbers).

- You are at least as uncertain about A_{odd,prime} = A_{odd} ∩ A_{prime} as you are about A_{odd}, simply because the later includes the former.
- You may attach $^{2}/_{6}$ to $A_{\text{odd,prime}}$ and $^{3}/_{6}$ to A_{odd} and to A_{prime} . You may have information about prime numbers being particularly rare (e.g., your dice was built that way), then you choose a qualifier for A_{prime} that's lower than that of A_{odd} , and an even lower qualifier for $A_{\text{odd,prime}}$.

- Your uncertainty is qualitatively different for different propositions.
- You attach an uncertainty qualifier to each proposition in Σ. A qualifier is anything you can order (e.g., fractions of a stick, numbers).
- (Ω, Σ) is a measurable space, hence you may use a plausibility measure to give a more detailed representation of your ignorance about propositions.

- You are at least as uncertain about A_{odd,prime} = A_{odd} ∩ A_{prime} as you are about A_{odd}, simply because the later includes the former.
- You may attach 2/6 to $A_{\text{odd,prime}}$ and 3/6 to A_{odd} and to A_{prime} . You may have information about prime numbers being particularly rare (e.g., your dice was built that way), then you choose a qualifier for A_{prime} that's lower than that of A_{odd} , and an even lower qualifier for $A_{\text{odd,prime}}$.
- Plausibility (Pl) generalises certain aspects of a typical measure (e.g., it may be non-numerical). You can now convey $Pl(A_{odd}) \ge Pl(A_{prime}) \ge Pl(A_{odd,prime})$, and much more.

- Your uncertainty is qualitatively different for different propositions.
- You attach an uncertainty qualifier to each proposition in Σ. A qualifier is anything you can order (e.g., fractions of a stick, numbers).
- (Ω, Σ) is a measurable space, hence you may use a plausibility measure to give a more detailed representation of your ignorance about propositions.
- Plausibility measures include belief functions (Shafer, 1976), possibility measures (Dubois and Prade, 1990), ordinal ranking functions (Goldszmidt and Pearl, 1992), (non-numerical) preference orders (Friedman and Halpern, 1996), and, of course, probability (Kolmogorov, 1960).

- You are at least as uncertain about A_{odd,prime} = A_{odd} ∩ A_{prime} as you are about A_{odd}, simply because the later includes the former.
- You may attach 2/6 to $A_{odd,prime}$ and 3/6 to A_{odd} and to A_{prime} . You may have information about prime numbers being particularly rare (e.g., your dice was built that way), then you choose a qualifier for A_{prime} that's lower than that of A_{odd} , and an even lower qualifier for $A_{odd,prime}$.
- Plausibility (Pl) generalises certain aspects of a typical measure (e.g., it may be non-numerical). You can now convey $Pl(A_{odd}) \ge Pl(A_{prime}) \ge Pl(A_{odd,prime})$, and much more.
- Under certain documented assumptions (Friedman and Halpern, 1996), they enable something like a 'calculus of uncertainty' which formalises the procedures the agent must follow to incorporate additional information about the world and revise their uncertainty representation coherently (in axiomatic probability, this is known as *conditioning*).

- Your uncertainty is qualitatively different for different propositions.
- You attach an uncertainty qualifier to each proposition in Σ. A qualifier is anything you can order (e.g., fractions of a stick, numbers).
- (Ω, Σ) is a measurable space, hence you may use a plausibility measure to give a more detailed representation of your ignorance about propositions.
- Plausibility measures include belief functions (Shafer, 1976), possibility measures (Dubois and Prade, 1990), ordinal ranking functions (Goldszmidt and Pearl, 1992), (non-numerical) preference orders (Friedman and Halpern, 1996), and, of course, probability (Kolmogorov, 1960).

Probability

The most well known plausibility measure.

Probability has been motivated from various angles, the most prominent:

- Objectivist: a notion of long-run stable frequency of repeatable events
- Subjectivist: a personal quantification of belief, it owes nothing to sample frequency (though it may coincide with it whenever that makes sense to an agent), it is constrained only by the axioms of probability theory and not by any interpretation (as chance or frequency).
 (Ramsey, 1931; De Finetti, 1974)

- Objectivist view is coherent with definition (ii) *the state of not being definitely known or perfectly clear;*
- Subjectivist view is coherent with definition (iii) the state or character of being uncertain in mind; a state of doubt

Both views share the same formal device (i.e., probability measures). The different interpretations are relevant when discussing strategies to use data to inform our representation of uncertainty.

Historical overview of different interpretations (Hacking, 1975)

Probability measure

A function $\mathsf{Pr}:\Sigma\to[0,1]$ such that

• $\Pr(\emptyset) = 0$

• $Pr(\Omega) = 1$

• and
$$\Pr(\bigcup_i A_i) = \sum_i \Pr(A_i)$$
 for pairwise disjoint events A_i
additivity

The significance of (countable) additivity is tremendous.

It can be shown (Radon–Nikodym theorem) that to identify a probability measure (i.e., over an event space Σ) it is sufficient to identify a probability density function (which assigns a non-negative density to each *outcome*, rather than event) and a base measure.

It is much easier to work with probability density functions than with probability measures directly, esp when we intend to *predict* these objects (e.g., using NNs) from available information. Pr is a function from $\Sigma \rightarrow [0, 1]$ s.t. countable additivity, a pdf is a function from \mathbb{R} to $\mathbb{R}_{>0}$ whose integral converges (to 1 if properly normalised).

Statistics

Gives us procedures we can use to fix the "free parameters" of our favourite framework of uncertainty representation (e.g., the probability measure) as to be coherent with relevant knowledge and evidence.

- Frequentist statistics: deeply rooted in the objectivist interpretation; procedures are based on repeatedly sampling data and typically formulated as optimisation problems (e.g., maximum likelihood estimation).
- Bayesian statistics: deeply rooted in subjectivist interpretation; procedures are based on probability calculus thus formulated as probabilistic inference problems (e.g., conditioning, marginalisation, expectation).

In ML

We represent our uncertainty about something by identifying a probability measure over a space of propositions (events) about the variables of interest; we typically specify a family of such measures and prioritise the members that are more consistent with observational data (that's the role of statistics).

If we use Frequentist procedures (which we often do, at least in DL), we are tempted to think that our representation will comply with the objectivist interpretation of probability. Unfortunately, this isn't really the case, but that is a topic for another chat :)

If you are developing a project on a topic such as calibration of NNs (esp calibration of LMs) then we *need* to have that chat. In the meantime, you will find these useful

- Baan et al. (2022): why you should not trust ECE under data uncertainty and what you can do instead; and what this means for LMs (Ilia and Aziz, 2024) and conditional generators (Giulianelli et al., 2023);
- for a more open-ended discussion (Baan et al., 2024)

Summary

- Uncertainty isn't variance, or entropy, or probability, or statistics, ...
- Uncertainty is a state of limited knowledge and it can be represented from the perspective of the phenomena under study or of the agents studying those.
- Possible worlds (a representation of what is possible) and plausibility measures (an expression of preferences) give a family of mathematical tools for uncertainty representation.
- Probability is the dominant tool in ML, partly due to its intuitive foundations (and interpretations), partly due to us having developed statistics for it, thus enabling data-driven procedures to represent uncertainty optimally (in some sense of the word).

Talk recommendation: for insight into the two mainstream views on uncertainty, watch Kristin Lennox's All About That Bayes https://youtu.be/eDMGDhyDxuY?si=skzXj7WC24Jc2nPt

Book recommendation: Lindley (2013) if you want to study uncertainty (from a probability standpoint) in great generality.

In the context of natural language generation, check (Baan et al., 2023).

Outline



Probabilistic Models

In Modelling Random Experiments

4 Modelling Observed Random Variables

5 Tools for prescribing distributions

- Univariate
- Multivariate

A probabilistic model *prescribes* the probability measure of a random experiment.

Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

A probabilistic model *prescribes* the probability measure of a random experiment.



Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

There are so many ways in which we can prescribe a probability measure:

• We may specify the probability of events in the event space, one at a time. This is at least tedious, intricate (we cannot violate countable additivity), sometimes impossible!

Figures from Wikimedia (CC-ASA-4.0)

A probabilistic model *prescribes* the probability measure of a random experiment.



Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

There are so many ways in which we can prescribe a probability measure:

- We may specify the probability of events in the event space, one at a time. This is at least tedious, intricate (we cannot violate countable additivity), sometimes impossible!
- We may instead specify a probability mass or density function (pmf or pdf) for outcomes of a random variable X : Ω → X ⊆ ℝ. The rv and its pdf f_X(x) in turn identify a probability measure via P(X ∈ A) = ∫_A f_X(x)dx.

Figures from Wikimedia (CC-ASA-4.0)

A probabilistic model *prescribes* the probability measure of a random experiment.



Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

There are so many ways in which we can prescribe a probability measure:

- We may specify the probability of events in the event space, one at a time. This is at least tedious, intricate (we cannot violate countable additivity), sometimes impossible!
- We may instead specify a probability mass or density function (pmf or pdf) for outcomes of a random variable X : Ω → X ⊆ ℝ. The rv and its pdf f_X(x) in turn identify a probability measure via P(X ∈ A) = ∫_A f_X(x)dx.
- We may instead specify the cumulative distribution function (cdf) of an rv, the cdf $F_X(x)$ in turn identifies a pdf via $f_X(x) = \frac{\mathrm{d}F_X(x)}{\mathrm{d}x}$, the rv and its pdf identify a probability measure.

Figures from Wikimedia (CC-ASA-4.0)

A probabilistic model *prescribes* the probability measure of a random experiment.



Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

There are so many ways in which we can prescribe a probability measure:

- We may specify the probability of events in the event space, one at a time. This is at least tedious, intricate (we cannot violate countable additivity), sometimes impossible!
- We may instead specify a probability mass or density function (pmf or pdf) for outcomes of a random variable X : Ω → X ⊆ ℝ. The rv and its pdf f_X(x) in turn identify a probability measure via P(X ∈ A) = ∫_A f_X(x)dx.
- We may instead specify the cumulative distribution function (cdf) of an rv, the cdf $F_X(x)$ in turn identifies a pdf via $f_X(x) = \frac{dF_X(x)}{dx}$, the rv and its pdf identify a probability measure.
- We may specify a simulator (e.g., a function from rand() to outcomes of an rv), the simulator identifies an inverse cdf F_X^{-1} , which in turn identifies a cdf, which in turn ...
Probabilistic Models

A probabilistic model *prescribes* the probability measure of a random experiment.



Random experiment: a sample space Ω , an event space $\Sigma = \mathcal{P}(\Omega)$, and a probability measure $\mathbb{P} : \Sigma \to [0,1]$ where $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$ and $\mathbb{P}(\bigcup_{i \in I} E_i) = \sum_{i \in I} \mathbb{P}(E_i)$ for collections $\{E_i\}$ of pairwise disjoint events.

There are so many ways in which we can prescribe a probability measure:

- We may specify the probability of events in the event space, one at a time. This is at least tedious, intricate (we cannot violate countable additivity), sometimes impossible!
- We may instead specify a probability mass or density function (pmf or pdf) for outcomes of a random variable X : Ω → X ⊆ ℝ. The rv and its pdf f_X(x) in turn identify a probability measure via P(X ∈ A) = ∫_A f_X(x)dx.
- We may instead specify the cumulative distribution function (cdf) of an rv, the cdf $F_X(x)$ in turn identifies a pdf via $f_X(x) = \frac{dF_X(x)}{dx}$, the rv and its pdf identify a probability measure.
- We may specify a simulator (e.g., a function from rand() to outcomes of an rv), the simulator identifies an inverse cdf F_X^{-1} , which in turn identifies a cdf, which in turn ...

Probabilistic Modelling and Reasoning

Probabilistic modelling concerns the specification of a joint distribution over random variables of interest.

Probabilistic reasoning concerns fixing a subset of these random variables to some observations and inferring marginal and conditional distributions by application of probability calculus.

The latter is also known as probabilistic inference.

Notation. Capital letters for rvs (e.g., X, Y), lowercase letters for assignments (e.g., X = x, Y = y), calligraphic letters for range of rvs (e.g., \mathcal{X} , \mathcal{Y}). I use p_X for the pdf of X and F_X for its cdf. When needed I show the dependency of the probability density on a parameter θ as follows: $p_X(x|\theta)$.

Probability calculus recap. Chain rule $p_{XY}(x, y) = p_X(x)p_{Y|X}(y|x) = p_Y(y)p_{X|Y}(x|y)$. Conditional probability $p_{Y|X}(y|x) = \frac{p_{XY}(x,y)}{p_X(x)}$. Marginalisation $p_X(x) = \int_{\mathcal{Y}} p_{XY}(x,y) \, dy$.

Oftentimes, we begin specifying a probability distribution by choosing a class of distributions (e.g., Normal, Exponential, Categorical).



Oftentimes, we begin specifying a probability distribution by choosing a class of distributions (e.g., Normal, Exponential, Categorical).

For multivariate data, we may need to choose a factorisation of the joint distribution.



Oftentimes, we begin specifying a probability distribution by choosing a class of distributions (e.g., Normal, Exponential, Categorical).

For multivariate data, we may need to choose a factorisation of the joint distribution.

Model fitting

- Maximum likelihood estimation (MLE) singles out a member of the class (e.g., $\mathcal{N}(2, 1)$, Exponential(10), Cat(0.1, 0.2, 0.7)).
- Bayesian estimation conditions on available evidence (data and model assumptions) to update prior beliefs (via probability calculus).





Oftentimes, we begin specifying a probability distribution by choosing a class of distributions (e.g., Normal, Exponential, Categorical).

For multivariate data, we may need to choose a factorisation of the joint distribution.

Model fitting

- Maximum likelihood estimation (MLE) singles out a member of the class (e.g., $\mathcal{N}(2, 1)$, Exponential(10), Cat(0.1, 0.2, 0.7)).
- Bayesian estimation conditions on available evidence (data and model assumptions) to update prior beliefs (via probability calculus).





Reasons for appreciating probabilistic models

Probabilistic models allow us to incorporate assumptions through

- the choice of distribution
- dependencies among random variables
- the way that distributions uses side information
- stipulate unobserved data and their properties

They return a distribution over outcomes which can be used to

- generate data
- account for unobserved data
- provide explanation and suggest improvements
- inform decision makers

17 / 66

Summary

A probabilistic model prescribes the probability measure of a random experiment.

Design: we can prescribe probability measures in many ways, some more or less implicit/indirect (and this will come in *very* handy).

Learning: we use data and statistics to fit the free parameters of our model.

Reasoning: we fix a subset of variables and perform marginal and/or conditional inferences.

Book recommendation: if you would like to learn *all* about probabilistic graphical models (PGMs), check the excellent book by Koller and Friedman (2009). I'd recommend Part I (on representation of distributions) to *anyone*.

Outline



2 Probabilistic Models

Modelling Random Experiments

4 Modelling Observed Random Variables

5 Tools for prescribing distributions

- Univariate
- Multivariate

Modelling random experiments

We treat data as outcomes of experiments involving random variables.

A *model* of the data prescribes a distribution for those random variables. Ideally, one that is faithful to statistical properties of our observations. Some applications:

- reveal structure hidden in existing data;
- support decisions about existing and future data.

The main subject of statistical interest is data (as opposed to tasks). Think of a task as a potential application of a (good) model of the data.

Modelling data does not imply solving a predictive task.

For example, a generative classifier is built upon a joint pdf $p_Y(y)p_{X|Y}(x|y)$ over labels $y \in \mathcal{Y}$ and inputs $x \in \mathcal{X}$. Making a specific prediction for a novel input x_* is a decision problem, oftentimes handled independently of model specification and learning.

A common decision rule for classification is

$$y_{*} = \arg \max_{c} p_{Y|X}(c|x_{*})$$
(1a)
$$= \arg \max_{c} \frac{p_{Y}(c)p_{X|Y}(x_{*}|c)}{p_{X}(x_{*})}$$
(1b)

A generalisation of it uses a utility function u(c, y) to compare a candidate c to an outcome of $Y|X = x_*$:

$$y_* = \arg\max_{c} \mathbb{E}_{p_{Y|X=x_*}}[u(c, Y)]$$
(2)

This also works for structure prediction, see an example of it in machine translation (Eikema and Aziz, 2022).

19/66



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)



- the measurements are continuous and positive
- the sample mean is close to 32
- the sample stddev is close to 16
- they concentrate around a single value (unimodal)
- they stretch to the right (skew)

Hidden structure



Here the measurements are natural numbers, the sample mean is close to 12.5 and the median is 12.

A Poisson distribution can capture the mean, but not the spread (recall that the Poisson mean and variance are equal).

Hidden structure



A different probabilistic model may posit the presence of two groups mixed in a single population.

Here the measurements are natural numbers, the sample mean is close to 12.5 and the median is 12.

A Poisson distribution can capture the mean, but not the spread (recall that the Poisson mean and variance are equal).

Decisions about future data



Here we observe continuous measurements from a sensor in a car. Data come in in batches of 100 measurements.

Suppose that if 1% (or more) of the readings drop below 0, the driver is at risk.

Decisions about future data



The heavier tails of the Student's t reserve much more probability for unseen data.

Here we observe continuous measurements from a sensor in a car. Data come in in batches of 100 measurements.

Suppose that if 1% (or more) of the readings drop below 0, the driver is at risk.

Summary

We treat *data* as outcomes of experiments involving random variables.

A model of the data prescribes a distribution for those random variables.

Rarely there is such a thing as a *correct* model. A model can be useful or mislead those using it. Ideally, a model

- is faithful to statistical properties of our observations
- reveals structure/patterns assumed to exist;
- supports decisions about existing and future data.

The main subject of statistical interest is data. A good model of the data has potential to power a task (e.g., a decision-making pipeline).

Book recommendation: Gelman et al. (2013), if you want to greatly enhance your statistical thinking. Andrew Gelman has made the book (and updates) freely available on his site.

Outline



2 Probabilistic Models

Modelling Random Experiments

4 Modelling Observed Random Variables

5 Tools for prescribing distributions

- Univariate
- Multivariate

Modelling observed random variables

Our goal is to learn a distribution over a set of **observed** random variables.

Observed random variables are the result of random experiments that have already happened: e.g., sentences in a collection of news articles, number of stars in a product review.

Modelling observed random variables

Our goal is to learn a distribution over a set of **observed** random variables.

Observed random variables are the result of random experiments that have already happened: e.g., sentences in a collection of news articles, number of stars in a product review.

Typical use in ML: conditional models.

▷ We are given some variables (inputs) and we are interested in making predictions about other variables (outputs)

- such inputs are also called *predictors* (or *covariates*)
- with some probability mass/density, *predicted by the model*, an output takes on a certain *outcome* in a sample space

| Predictor | Outcome | Sample space |
|--|---|--------------------|
| Why did they bother recording this??? | * | 1–5 stars |
| Source: geen standaard MT: no standard | compare('no step')=0.5 | [0, 1] |
| he proposed a famous solution to an inverse probability problem in the 18th century | https://en.wikipedia.org/wiki/ Thomas_Bayes | \mathcal{W}_{en} |
| who is the main mind behind what we call the Bayesian inter- pretation of probability? | You might be talking about Bayesian statistics, though it's named after Thomas Bayes (1701–1761), it was Pierre-Simon Laplace (1749–1827) who developed most of the theory. | Σ_{en}^{*} |
| | | |
| | Pepper loves the beach! | Σ_{en}^{*} |
| That's not possible! | Dat is niet mogelijk! | $\sum_{n=1}^{*}$ |

24 / 66

Modelling observed random variables

Our goal is to learn a distribution over a set of **observed** random variables.

Observed random variables are the result of random experiments that have already happened: e.g., sentences in a collection of news articles, number of stars in a product review.

Typical use in ML: conditional models.

▷ We are given some variables (inputs) and we are interested in making predictions about other variables (outputs)

- such inputs are also called *predictors* (or *covariates*)
- with some probability mass/density, *predicted by the model*, an output takes on a certain *outcome* in a sample space

| Predictor | Outcome | Sample space |
|--|---|--------------------|
| Why did they bother recording this??? | * | 1–5 stars |
| Source: geen standaard MT: no standard | compare('no step')=0.5 | [0, 1] |
| he proposed a famous solution to an inverse probability problem in the 18th century | https://en.wikipedia.org/wiki/ Thomas_Bayes | \mathcal{W}_{en} |
| who is the main mind behind what we call the Bayesian inter- pretation of probability? | You might be talking about Bayesian statistics, though it's named after Thomas Bayes (1701–1761), it was Pierre-Simon Laplace (1749–1827) who developed most of the theory. | Σ_{en}^{*} |
| | | |
| | Pepper loves the beach! | Σ_{en}^{*} |
| That's not possible! | Dat is niet mogelijk! | $\sum_{n=1}^{*}$ |

24 / 66

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



• Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally
- Do you expect variance to differ?

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally
- Do you expect variance to differ?
- Do you expect skew?

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally
- Do you expect variance to differ?
- Do you expect skew?
- Bounded support? Multimodality? Asymmetry?

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails



- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally
- Do you expect variance to differ?
- Do you expect skew?
- Bounded support? Multimodality? Asymmetry?
- Unsure about categorical or ordinal treatment?

- data type: countable (binary, categorical, ordinal), uncountable, univariate or multivariate, combinatorial, etc.
- e match properties of the data and distribution: overdispersion, skewness, heavy tails

- Don't be mislead by data sparsity, collect more measurements and you will observe uncertainty (Plank, 2022).
- Don't be mislead by the marginal if you intend to model conditionally
- Do you expect variance to differ?
- Do you expect skew?
- Bounded support? Multimodality? Asymmetry?
- Unsure about categorical or ordinal treatment?

Statistical models parameterised by NNs

Once a family of distributions is in place, we let a neural network predict a member of the family, which it does by mapping from available information (e.g., x).

 $Y|x \sim \mathsf{Cat}(f(x; \theta))$ or $Y|x \sim \mathcal{N}(\mu(x; \theta), \sigma(x; \theta)^2)$

we then proceed to estimate parameters θ of the NNs

 NNs compute the parameters of the statistical model. We estimate NN parameters.

Example - Text classifier

Before DL was popular, we would identifying informative features h(x) of the available predictor x. We would then map these features to the parameter of a Categorical distribution (e.g., via a log-linear model): $Y|X = x \sim \text{Cat}(\text{softmax}(Wh(x) + b)).$



Example - Text classifier

Before DL was popular, we would identifying informative features h(x) of the available predictor x. We would then map these features to the parameter of a Categorical distribution (e.g., via a log-linear model): $Y|X = x \sim \text{Cat}(\text{softmax}(Wh(x) + b)).$

Nowadays, we tend to condition on **everything available to us** by learning how to map from **arbitrarily complex** data to the parameters of our distributions. We do so with NNs: $Y|X = x \sim Cat(f(x; \theta))$.

There's a lot of research on how to design $f(\cdot; \theta)$ and estimate θ effectively.

In $Y|X = x \sim \text{Cat}(f(x; \theta))$, $f(\cdot; \theta)$ is a NN architecture with parameters θ , it maps any covariate x, say a long review in English, to the parameters of the Categorical distribution that *by assumption* govern the conditional response variable.


Example - Text classifier

Before DL was popular, we would identifying informative features h(x) of the available predictor x. We would then map these features to the parameter of a Categorical distribution (e.g., via a log-linear model): $Y|X = x \sim \text{Cat}(\text{softmax}(Wh(x) + b)).$

Nowadays, we tend to condition on **everything available to us** by learning how to map from **arbitrarily complex** data to the parameters of our distributions. We do so with NNs: $Y|X = x \sim Cat(f(x; \theta))$.

There's a lot of research on how to design $f(\cdot; \theta)$ and estimate θ effectively.

In $Y|X = x \sim \text{Cat}(f(x; \theta))$, $f(\cdot; \theta)$ is a NN architecture with parameters θ , it maps any covariate x, say a long review in English, to the parameters of the Categorical distribution that *by assumption* govern the conditional response variable.



We have a probability model of a random variable Y, and this model may condition on available covariates X. This model has parameters θ and assigns probability mass/density $p(y|x, \theta)$ to an observation.

I may omit the subscripts from the pdfs whenever I find it unambiguous. That is, I write $p(y|x, \theta)$ instead of $p_{Y|X}(y|x, \theta)$.

We have a probability model of a random variable Y, and this model may condition on available covariates X. This model has parameters θ and assigns probability mass/density $p(y|x, \theta)$ to an observation.

Given a dataset $\mathcal{D} = \{(x^{(1)}, y^{(1)}), \dots, (x^{(N)}, y^{(N)})\}$ of i.i.d. observations,

I may omit the subscripts from the pdfs whenever I find it unambiguous. That is, I write $p(y|x, \theta)$ instead of $p_{Y|X}(y|x, \theta)$.

We have a probability model of a random variable Y, and this model may condition on available covariates X. This model has parameters θ and assigns probability mass/density $p(y|x, \theta)$ to an observation.

Given a dataset $\mathcal{D} = \{(x^{(1)}, y^{(1)}), \dots, (x^{(N)}, y^{(N)})\}$ of i.i.d. observations, the log-likelihood function gives us a criterion for parameter estimation

 $\mathcal{L}_{\mathcal{D}}(\theta) =$

I may omit the subscripts from the pdfs whenever I find it unambiguous. That is, I write $p(y|x, \theta)$ instead of $p_{Y|X}(y|x, \theta)$.

We have a probability model of a random variable Y, and this model may condition on available covariates X. This model has parameters θ and assigns probability mass/density $p(y|x, \theta)$ to an observation.

Given a dataset $\mathcal{D} = \{(x^{(1)}, y^{(1)}), \dots, (x^{(N)}, y^{(N)})\}$ of i.i.d. observations, the log-likelihood function gives us a criterion for parameter estimation

$$\mathcal{L}_{\mathcal{D}}(heta) = \log \prod_{s=1}^{N} p(y^{(s)}|x^{(s)}, heta) =$$

I may omit the subscripts from the pdfs whenever I find it unambiguous. That is, I write $p(y|x, \theta)$ instead of $p_{Y|X}(y|x, \theta)$.

We have a probability model of a random variable Y, and this model may condition on available covariates X. This model has parameters θ and assigns probability mass/density $p(y|x, \theta)$ to an observation.

Given a dataset $\mathcal{D} = \{(x^{(1)}, y^{(1)}), \dots, (x^{(N)}, y^{(N)})\}$ of i.i.d. observations, the log-likelihood function gives us a criterion for parameter estimation

I may omit the subscripts from the pdfs whenever I find it unambiguous. That is, I write $p(y|x, \theta)$ instead of $p_{Y|X}(y|x, \theta)$.

$$\mathcal{L}_{\mathcal{D}}(\theta) = \log \prod_{s=1}^{N} p(y^{(s)} | x^{(s)}, \theta) = \sum_{s=1}^{N} \log p(y^{(s)} | x^{(s)}, \theta)$$

If the log-likelihood is **differentiable** and **tractable** then backpropagation gives us the gradient

 $oldsymbol{
abla}_{ heta}\mathcal{L}_{\mathcal{D}}(heta) =$

Differentiable

Consider the example of a Categorical likelihood:

- for a data point (x, y) the log-likelihood is log Cat(y|f(x; θ)) = log f_y(x; θ) This shows that the Categorical likelihood Cat(y|f(x; θ)) is differentiable with respect to its parameter f_y(x; θ).
- To satisfy differentiability with respect to θ for any (x, y), we need f(·; θ), to be differentiable with respect to θ in its domain (the space X of all covariates).

Tractable The evaluation of $f(x; \theta)$ is tractable for any $x \in \mathcal{X}$.

Beyond Think about other likelihoods (e.g., Bernoulli, Binomial, Multinomial, Poisson, Geometric, Gaussian, Exponential, Gamma), can you imagine differentiable and tractable parameterisations of the model?

If the log-likelihood is **differentiable** and **tractable** then backpropagation gives us the gradient

$$oldsymbol{
abla}_{ heta}\mathcal{L}_{\mathcal{D}}(heta) = oldsymbol{
abla}_{ heta}\sum_{s=1}^{N}\log p(y^{(s)}|x^{(s)}, heta) =$$

Differentiable

Consider the example of a Categorical likelihood:

- for a data point (x, y) the log-likelihood is log Cat(y|f(x; θ)) = log f_y(x; θ) This shows that the Categorical likelihood Cat(y|f(x; θ)) is differentiable with respect to its parameter f_y(x; θ).
- To satisfy differentiability with respect to θ for any (x, y), we need f(·; θ), to be differentiable with respect to θ in its domain (the space X of all covariates).

Tractable The evaluation of $f(x; \theta)$ is tractable for any $x \in \mathcal{X}$.

Beyond Think about other likelihoods (e.g., Bernoulli, Binomial, Multinomial, Poisson, Geometric, Gaussian, Exponential, Gamma), can you imagine differentiable and tractable parameterisations of the model?

If the log-likelihood is **differentiable** and **tractable** then backpropagation gives us the gradient

$$oldsymbol{
abla}_{ heta}\mathcal{L}_{\mathcal{D}}(heta) = oldsymbol{
abla}_{ heta}\sum_{s=1}^{N}\log p(y^{(s)}|x^{(s)}, heta) = \sum_{s=1}^{N}oldsymbol{
abla}_{ heta}\log p(y^{(s)}|x^{(s)}, heta)$$

Differentiable

Consider the example of a Categorical likelihood:

- for a data point (x, y) the log-likelihood is log Cat(y|f(x; θ)) = log f_y(x; θ) This shows that the Categorical likelihood Cat(y|f(x; θ)) is differentiable with respect to its parameter f_y(x; θ).
- To satisfy differentiability with respect to θ for any (x, y), we need f(·; θ), to be differentiable with respect to θ in its domain (the space X of all covariates).

Tractable The evaluation of $f(x; \theta)$ is tractable for any $x \in \mathcal{X}$.

Beyond Think about other likelihoods (e.g., Bernoulli, Binomial, Multinomial, Poisson, Geometric, Gaussian, Exponential, Gamma), can you imagine differentiable and tractable parameterisations of the model?

If the log-likelihood is **differentiable** and **tractable** then backpropagation gives us the gradient

$$oldsymbol{
abla}_{ heta}\mathcal{L}_{\mathcal{D}}(heta) = oldsymbol{
abla}_{ heta}\sum_{s=1}^{N}\log p(y^{(s)}|x^{(s)}, heta) ~~=~ \sum_{s=1}^{N}oldsymbol{
abla}_{ heta}\log p(y^{(s)}|x^{(s)}, heta)$$

and we can update $\boldsymbol{\theta}$ in the direction

 $\gamma \nabla_{\theta} \mathcal{L}_{\mathcal{D}}(\theta)$

to attain a local maximum of the likelihood function

Differentiable

Consider the example of a Categorical likelihood:

- for a data point (x, y) the log-likelihood is log Cat(y|f(x; θ)) = log f_y(x; θ) This shows that the Categorical likelihood Cat(y|f(x; θ)) is differentiable with respect to its parameter f_y(x; θ).
- To satisfy differentiability with respect to θ for any (x, y), we need f(·; θ), to be differentiable with respect to θ in its domain (the space X of all covariates).

Tractable The evaluation of $f(x; \theta)$ is tractable for any $x \in \mathcal{X}$.

Beyond Think about other likelihoods (e.g., Bernoulli, Binomial, Multinomial, Poisson, Geometric, Gaussian, Exponential, Gamma), can you imagine differentiable and tractable parameterisations of the model?

For large N, computing the gradient is inconvenient

$$\nabla_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \underbrace{\sum_{s=1}^{N} \nabla_{\theta} \log p(y^{(s)} | x^{(s)}, \theta)}_{\text{too many terms}}$$

We are looking for a principled way to approximate the exact gradient. Being principled here means enjoying some guarantees (this usually requires satisfying certain properties, as we shall see).

Note that we introduced the notion of a *stochastic gradient*, a random variable whose range is the space of gradient vectors of our model's log-likelihood function.

For large N, computing the gradient is inconvenient



We are looking for a principled way to approximate the exact gradient. Being principled here means enjoying some guarantees (this usually requires satisfying certain properties, as we shall see).

Note that we introduced the notion of a *stochastic gradient*, a random variable whose range is the space of gradient vectors of our model's log-likelihood function.

For large N, computing the gradient is inconvenient



We are looking for a principled way to approximate the exact gradient. Being principled here means enjoying some guarantees (this usually requires satisfying certain properties, as we shall see).

Note that we introduced the notion of a *stochastic gradient*, a random variable whose range is the space of gradient vectors of our model's log-likelihood function.

For large N, computing the gradient is inconvenient

$$\nabla_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \underbrace{\sum_{s=1}^{N} \nabla_{\theta} \log p(y^{(s)} | x^{(s)}, \theta)}_{\text{too many terms}}$$
$$= \sum_{s=1}^{N} \frac{1}{N} N \nabla_{\theta} \log p(y^{(s)} | x^{(s)}, \theta)$$
$$= \sum_{s=1}^{N} \mathcal{U}(s | 1/N) N \nabla_{\theta} \log p(y^{(s)} | x^{(s)}, \theta)$$
$$= \mathbb{E}_{S \sim \mathcal{U}(1/N)} \left[N \nabla_{\theta} \log p(y^{(S)} | x^{(S)}, \theta) \right]$$

S selects data points uniformly at random

We are looking for a principled way to approximate the exact gradient. Being principled here means enjoying some guarantees (this usually requires satisfying certain properties, as we shall see).

Note that we introduced the notion of a *stochastic gradient*, a random variable whose range is the space of gradient vectors of our model's log-likelihood function.

For large N, we can use a gradient estimate

 $\boldsymbol{\nabla}_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \mathbb{E}_{S \sim \mathcal{U}(1/N)} \left[N \boldsymbol{\nabla}_{\theta} \log p(\boldsymbol{y}^{(S)} | \boldsymbol{x}^{(S)}, \theta) \right]$ expected gradient :)

The theory of stochastic optimisation (Robbins and Monro, 1951) tells us that we will converge to a local optimum of the objective as long as we take steps that are correct *on average*. This means we can optimise with stochastic gradient estimates, for as long as they are unbiased estimates of the exact gradient.

Do you see the guarantee and the condition?

There are more conditions, however. The learning rate must comply with some key properties. Luckily many learning rate schedules have been documented in the literature, and most our famous optimisers meet the Robbis and Monro conditions (though not all).

If you want to read more, but need something more accessible than the 1951 paper, check (Bottou, 2010).

For large N, we can use a gradient estimate

$$\nabla_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \underbrace{\mathbb{E}_{S \sim \mathcal{U}(1/N)} \left[N \nabla_{\theta} \log p(y^{(S)} | x^{(S)}, \theta) \right]}_{\text{expected gradient :)}}$$
$$\stackrel{\text{MC}}{\approx} \frac{1}{M} \sum_{m=1}^{M} N \nabla_{\theta} \log p(y^{(s_m)} | x^{(s_m)}, \theta) \quad \text{with } S_m \sim \mathcal{U}(1/N)$$

The theory of stochastic optimisation (Robbins and Monro, 1951) tells us that we will converge to a local optimum of the objective as long as we take steps that are correct *on average*. This means we can optimise with stochastic gradient estimates, for as long as they are unbiased estimates of the exact gradient.

Do you see the guarantee and the condition?

There are more conditions, however. The learning rate must comply with some key properties. Luckily many learning rate schedules have been documented in the literature, and most our famous optimisers meet the Robbis and Monro conditions (though not all).

If you want to read more, but need something more accessible than the 1951 paper, check (Bottou, 2010).

For large N, we can use a gradient estimate $\boldsymbol{\nabla}_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \underbrace{\mathbb{E}_{S \sim \mathcal{U}(^{1}/N)} \left[N \boldsymbol{\nabla}_{\theta} \log p(y^{(S)} | x^{(S)}, \theta) \right]}_{\boldsymbol{\mathcal{U}}}$ expected gradient :) $\overset{\text{MC}}{\approx} \frac{1}{M} \sum_{m=1}^{M} N \boldsymbol{\nabla}_{\theta} \log p(\boldsymbol{y}^{(s_m)} | \boldsymbol{x}^{(s_m)}, \theta) \quad \text{with } S_m \sim \mathcal{U}(1/N)$ $= \boldsymbol{\nabla}_{\theta} \frac{N}{M} \sum_{m=1}^{M} \log p(\boldsymbol{y}^{(s_m)} | \boldsymbol{x}^{(s_m)}, \theta)$ $\mathcal{L}_{\mathcal{B}}(\theta)$

The theory of stochastic optimisation (Robbins and Monro, 1951) tells us that we will converge to a local optimum of the objective as long as we take steps that are correct *on average*. This means we can optimise with stochastic gradient estimates, for as long as they are unbiased estimates of the exact gradient.

Do you see the guarantee and the condition?

There are more conditions, however. The learning rate must comply with some key properties. Luckily many learning rate schedules have been documented in the literature, and most our famous optimisers meet the Robbis and Monro conditions (though not all).

If you want to read more, but need something more accessible than the 1951 paper, check (Bottou, 2010).

For large N, we can use a gradient estimate $\boldsymbol{\nabla}_{\theta} \mathcal{L}_{\mathcal{D}}(\theta) = \underbrace{\mathbb{E}_{S \sim \mathcal{U}(1/N)} \left[N \boldsymbol{\nabla}_{\theta} \log p(\boldsymbol{y}^{(S)} | \boldsymbol{x}^{(S)}, \theta) \right]}_{\boldsymbol{\mathcal{V}}_{\theta}}$ expected gradient :) $\overset{\text{MC}}{\approx} \frac{1}{M} \sum_{m=1}^{M} N \boldsymbol{\nabla}_{\theta} \log p(\boldsymbol{y}^{(s_m)} | \boldsymbol{x}^{(s_m)}, \theta) \quad \text{with } S_m \sim \mathcal{U}(1/N)$ $= \nabla_{\theta} \frac{N}{M} \sum_{m=1}^{M} \log p(y^{(s_m)} | x^{(s_m)}, \theta)$ $\mathcal{L}_{\mathcal{B}}(\theta)$

and take a step in the direction

 $\gamma \frac{N}{M} \underbrace{\nabla_{\theta} \mathcal{L}_{\mathcal{B}}(\theta)}_{\text{stochastic gradient}}$

The theory of stochastic optimisation (Robbins and Monro, 1951) tells us that we will converge to a local optimum of the objective as long as we take steps that are correct *on average*. This means we can optimise with stochastic gradient estimates, for as long as they are unbiased estimates of the exact gradient.

Do you see the guarantee and the condition?

There are more conditions, however. The learning rate must comply with some key properties. Luckily many learning rate schedules have been documented in the literature, and most our famous optimisers meet the Robbis and Monro conditions (though not all).

If you want to read more, but need something more accessible than the 1951 paper, check (Bottou, 2010).

where $\mathcal{B} = \{(x^{(s_1)}, y^{(s_1)}), \dots, (x^{(s_M)}, y^{(s_M)})\}$ is a random mini-batch

Summary - a recipe for supervised learning

Maximum likelihood estimation

- tells you which loss to optimise (i.e. negative log-likelihood)
- Automatic differentiation (backprop) with gradient surrogates
 - a tractable and differentiable forward computation whose backward is an unbiased estimate of the intended gradient
- Stochastic optimisation powered by backprop
 - general purpose gradient-based optimisers

Our main job is to pick an appropriate family of distributions.

Paper recommendation: for a comprehensive understanding of stochastic computation graphs (Schulman et al., 2015).

Outline



2 Probabilistic Models

Modelling Random Experiments

4 Modelling Observed Random Variables

(5) Tools for prescribing distributions

- Univariate
- Multivariate

Prescribing distributions

We will now discuss various ways to prescribe distributions using deep learning. For each technique, we will keep an eye on two things:

• our ability to assess the probability mass/density of a given outcome

• our ability to sample outcomes from the corresponding distribution We begin with the univariate case and then discuss the multivariate case.

- assessment: useful for learning (e.g., via approximate MLE).
- sampling: useful for making predictions.

Outline



2 Probabilistic Models

In Modelling Random Experiments

4 Modelling Observed Random Variables

5 Tools for prescribing distributions• Univariate

Multivariate

Overview

- enumeration
- known parametric form
- transform a known random source
- data augmentation and marginalisation

Throughout we parameterise the conditional distribution of a random variable Y given X = x.

We use neural network architectures which we denote with a shorthand notation:

layer_O(inputs; parameters)

e.g., linear_K(\mathbf{h} ; θ_{out}) uses parameters θ_{out} to map some input vector \mathbf{h} to K outputs linearly and deterministically.

We omit the implementation details hoping the layer's name is suggestive enough (e.g., linear is implemented as Wh + b with $\theta_{out} = \{W, b\}$).

Enumerate masses

Predict the probability mass of each and every one of K outcomes.



- mass assessment: evaluate K masses (e.g., an NN forward pass, t_{Δ}) and look the relevant one up
- sampling: linear in K (via inverse cdf or Gumbel trick)

Sentiment classifier

- Input: a piece of text x
- Output: a distribution over 3 possible sentiment labels (negative, neutral, positive).

Examples

- encode the text $\mathbf{h} = \text{encode}_D(x; \theta_{\text{enc}})$
- predict K scores $\mathbf{s} = (s_1, \ldots, s_K)^\top$
 - $\mathbf{s} = \mathsf{linear}_{\mathcal{K}}(\mathbf{h}; \theta_{\mathsf{out}})$
- map them to the probability simplex: $\pmb{\pi} = {\sf t}_\Delta({\sf s})$
 - softmax (see (Niculae and Blondel, 2017) and (Niculae et al., 2023, §3.2 and 3.3) for an origin story)
 - or sparsemax (Martins and Astudillo, 2016)
 - or entmax $_{\alpha}$ (Peters et al., 2019)
 - etc.

Enumerate masses

Predict the probability mass of each and every one of K outcomes.



- mass assessment: evaluate K masses (e.g., an NN forward pass, t_{Δ}) and look the relevant one up
- sampling: linear in K (via inverse cdf or Gumbel trick)

How about countably infinite spaces?

Sentiment classifier

- Input: a piece of text x
- Output: a distribution over 3 possible sentiment labels (negative, neutral, positive).

Examples

- encode the text $\mathbf{h} = \text{encode}_D(x; \theta_{enc})$
- predict K scores $\mathbf{s} = (s_1, \ldots, s_K)^\top$
 - $\mathbf{s} = \text{linear}_{K}(\mathbf{h}; \theta_{\text{out}})$
- map them to the probability simplex: $\pmb{\pi} = t_\Delta(\pmb{s})$
 - softmax (see (Niculae and Blondel, 2017) and (Niculae et al., 2023, §3.2 and 3.3) for an origin story)
 - or sparsemax (Martins and Astudillo, 2016)
 - or entmax $_{\alpha}$ (Peters et al., 2019)
 - etc.

Known pmf

Predict the parameter(s) of a known pmf.



- mass assessment: evaluate the parameter(s) (e.g., NN forward pass) and the pmf (a few operations)
- sampling is typically possible (via inverse cdf method, or some specialised algorithm)

Heard count

- Input: an image $\mathbf{x} \in \mathbb{R}^{H \times W \times C}$ of a field
- Output: a Poisson distribution over the number of cows in the field. A Poisson is identified by its rate parameter (a strictly positive scalar), which we predicted given **x** as follows

$$\begin{aligned} \mathbf{h} &= \operatorname{encode}_{D}(\mathbf{x}; \theta_{\mathsf{enc}}) \\ \lambda &= \operatorname{softplus}(\operatorname{linear}_{1}(\mathbf{h}; \theta_{\mathsf{rate}})) \end{aligned}$$
(3)

Probability mass of an outcome: Poisson $(k|\lambda) = \frac{\lambda^k \exp(-k)}{k!}$

Inverse cdf method: if $icdf(\cdot|\lambda)$ is the inverse of the cdf of the rv Y, then transforming a uniform sample $p \sim \mathcal{U}(0,1)$ via $icdf(p|\lambda)$ yields an rv with the same distribution as Y.

Known pmf

Predict the parameter(s) of a known pmf.



- mass assessment: evaluate the parameter(s) (e.g., NN forward pass) and the pmf (a few operations)
- sampling is typically possible (via inverse cdf method, or some specialised algorithm)

How about uncountable sample spaces?

Deep Learning 2 @ UvA

Representing Uncertainty in ML

Heard count

- Input: an image $\mathbf{x} \in \mathbb{R}^{H \times W \times C}$ of a field
- Output: a Poisson distribution over the number of cows in the field. A Poisson is identified by its rate parameter (a strictly positive scalar), which we predicted given **x** as follows

$$\begin{aligned} \mathbf{h} &= \operatorname{encode}_{D}(\mathbf{x}; \theta_{\mathsf{enc}}) \\ \lambda &= \operatorname{softplus}(\operatorname{linear}_{1}(\mathbf{h}; \theta_{\mathsf{rate}})) \end{aligned}$$
 (4)

Probability mass of an outcome: Poisson $(k|\lambda) = \frac{\lambda^k \exp(-k)}{k!}$

Inverse cdf method: if $icdf(\cdot|\lambda)$ is the inverse of the cdf of the rv Y, then transforming a uniform sample $p \sim \mathcal{U}(0,1)$ via $icdf(p|\lambda)$ yields an rv with the same distribution as Y.

Known pdf

Predict the parameter(s) of a known pdf.



- density assessment: evaluate the parameter(s) (e.g., NN forward pass) and the pdf (a few operations, assuming analytical form or an efficient numerical algorithm)
- sampling is typically possible (via inverse cdf method, or some specialised algorithm)

Temperature

- Input: an image $\mathbf{x} \in \mathbb{R}^{H \times W \times C}$ of a car engine
- Output: a Normal distribution over temperature values in Celsius. A Normal is identified by a location in \mathbb{R} and a scale in $\mathbb{R}_{>0}$, which we predict given x as follows

$$\mathbf{h} = \text{encode}_{D}(\mathbf{x}; \theta_{\text{enc}})$$

$$\mu = \text{linear}_{1}(\mathbf{h}; \theta_{\text{loc}})$$

$$\sigma = \text{softplus}(\text{linear}_{1}(\mathbf{h}; \theta_{\text{scale}}))$$
(5)

Probability density of an outcome: $\mathcal{N}(y|\mu, \sigma^2) = \frac{\exp\left(\frac{-(y-\mu)^2}{2\sigma^2}\right)}{\sigma\sqrt{2\pi}}$

Known pdf

Predict the parameter(s) of a known pdf.



- density assessment: evaluate the parameter(s) (e.g., NN forward pass) and the pdf (a few operations, assuming analytical form or an efficient numerical algorithm)
- sampling is typically possible (via inverse cdf method, or some specialised algorithm)

What if not flexible enough?

Deep Learning 2 @ UvA

Representing Uncertainty in ML

Temperature

- Input: an image $\mathbf{x} \in \mathbb{R}^{H \times W \times C}$ of a car engine
- Output: a Normal distribution over temperature values in Celsius. A Normal is identified by a location in ℝ and a scale in ℝ_{>0}, which we predict given x as follows

$$\mathbf{h} = \text{encode}_{D}(\mathbf{x}; \theta_{\text{enc}})$$

$$\mu = \text{linear}_{1}(\mathbf{h}; \theta_{\text{loc}}) \qquad (6)$$

$$\sigma = \text{softplus}(\text{linear}_{1}(\mathbf{h}; \theta_{\text{scale}}))$$

Probability density of an outcome: $\mathcal{N}(y|\mu, \sigma^2) = \frac{\exp\left(\frac{-(y-\mu)^2}{2\sigma^2}\right)}{\sigma\sqrt{2\pi}}$

Unnormalised pdf/pmf

Predict a non-negative score for a *given* outcome (as opposed to each and every outcome). This procedure must identify a pdf up to an unknown normalisation constant.



- density assessment: difficult, numerical integration is possible (but inefficient in high dimensions)
- sampling: difficult, requires MC (inefficient in high dimensions) or MCMC (often slow, asymptotically correct).

We need to output non-negative values all over the domain, and need the integral across the domain to converge to a real number (that is, finite).

Here is an example, engineered to meet the requirements: a conic combination of exponentiated-square basis functions.

$$\begin{split} \mathbf{h} &= \operatorname{encode}_{D}(\mathbf{x}; \theta_{\operatorname{enc}}) \\ \mathbf{u} &= \operatorname{linear}_{C}(\mathbf{x}; \theta_{\operatorname{hid}}) \\ \mathbf{v} &= \exp(-\mathbf{u} \odot \mathbf{u}) & \leftarrow C \text{ non-negative numbers} \\ s &= \sum_{c=1}^{C} \exp(w_{c})v_{c} & \leftarrow \text{ non-negative slopes} \end{split}$$
 (7)

For countably infinite or uncountable sample spaces, it may be difficult to guarantee that the integral of an arbitrary non-negative function converges.

Energy-based models (EBMs), which we will cover in this module, are examples of such tools.

Deep Learning 2 @ UvA

Representing Uncertainty in ML

CDF

Predict the cumulative probability p for an outcome y: that is $p = \int_{a \le y} pdf(a|\theta) da$



- density assessment: requires differentiation (which can be automated!)
- sampling: difficult (requires inverting the cdf)

We can constrain an NN to specify a non-decreasing function by having non-negative weights in linear layers (biases are unconstrained), and non-decreasing activation functions (e.g., tanh, relu, softplus, sigmoid, etc.).

To constrain the output to [0, 1] we may use a sigmoid output, or a convex combination of *C* sigmoid outputs.

See, for example, (Huang et al., 2018; De Cao et al., 2020)

Percentile function - inverse cdf

Predict the outcome y for a given percentile p:



- density assessment: possible in some cases (via inverse function theorem)
- sampling: easy by design

Now the output is an outcome (if $\mathcal{Y} = \mathbb{R}$, the output is unconstrained), the domain of the NN is [0, 1], and the function must be non-decreasing (as the cdf case).

For y whose inverse p we know (for example, those we sampled via $p \sim U(0,1)$ and $y = icdf(p|\theta)$) we can assess the density using autodiff.

Sampler - bijection

Parameterise a bijective transformation of a known and convenient base random variable

- Assessment and sampling can be made simple, in some cases only one of the two is simple;
- Parameterising one such model takes some skill (to achieve efficient computations)

Suppose access to an invertible function (a bijection): $x = h^{-1}(y)$, then

$$p_Y(y) = p_X(h^{-1}(y)) |\det J_{h^{-1}}(y)|$$
(8)

Start from a known p_X , for example a Gaussian, and obtain a novel p_Y , more complex than a Gaussian.

It's possible to assess the density of some outcome y by mapping it to the corresponding $x = h^{-1}(y)$, assessing its density and the Jacobian.

It's possible to draw samples, by drawing from the simple p_X and then mapping to the corresponding y = h(x).

Normalising Flows (Rezende and Mohamed, 2015; Kingma et al., 2016; Papamakarios et al., 2019), which we cover in this module, are an example of such tool.

Sampler - general case

Parameterise an arbitrary transformation of a base random variable

- Sampling: trivial by design (it costs a forward pass through an NN we choose)
- Assessment: intractable in general! For an outcome y ∈ ℝ^O, a base density p_X(x) on ℝ^I

Here
$$f$$
 is some deep neural network with I inputs and O outputs (e.g., a feed forward neural network)

$$x \sim \mathcal{N}(0, I)$$
 (11)

$$y = f(x;\theta) \tag{12}$$

$$p_{Y}(y) = \int_{\mathcal{C}} p_{X}(x) dx$$
(9)
$$\mathcal{C} = \{x : f(x; \theta) = y\}$$
(10)

C is the set of all points in \mathbb{R}^{I} that f maps to exactly y

See Generative Adversarial Network (GAN; Goodfellow et al., 2014) and implicit distributions (Huszár, 2017).

Marginalise a latent variable - mixture model

Predict the parameters of *C* known pdfs and the coefficients **w** of a finite mixture: $p(y|x, \theta) = \sum_{z=1}^{C} w_z p(y|x, \theta_z)$ with $\mathbf{w} \in \Delta_{C-1}$.



This simple idea can be used to create very flexible distributions. See, for example, Farinhas et al. (2022), where the components are defined in different subsets of Δ .

- assessment: linear in C
- sampling: ancestral sampling (draw a component in linear time, then draw from that component)

Deep Learning 2 @ UvA

Marginalise a latent variable - compounding

Predict a distribution for the parameter(s) of a known parametric family: e.g., $\int_{\mathbb{R}_{>0}} \operatorname{Gamma}(\lambda | \alpha, \beta) \operatorname{Poisson}(y | \lambda) d\lambda$.



assessment: typically intractable (some exceptions in the exponential family), a common solution is a variational lowerbound
 commission exception

sampling: ancestral sampling

Deep Learning 2 @ UvA

Representing Uncertainty in ML

Where the Gamma's shape α and rate β are predicted by an NN.

Variation: have λ be predicted by an NN that takes $\mathbf{x}, \alpha, \beta$ as inputs.

- assessment: intractable
- sampling: ancestral sampling

Variational auto-encoders (Rezende et al., 2014; Kingma and Welling, 2014), which we cover in this module, are an example of such tool.
Outline



2 Probabilistic Models

Modelling Random Experiments

4 Modelling Observed Random Variables

5 Tools for prescribing distributions• Univariate

• Multivariate

Multivariate

For the fixed-dimension case, we may have access to multivariate generalisations of known pdfs (e.g., MVN).

In general (fixed-dimension or not), we can exploit a *factorisation* with or without conditional independence assumptions.

Then, we predict the factors:

- direct, or cdf, or sampler/simulator
- unnormalised

Directed

Decompose an outcome into parts $y = (w_1, \ldots, w_N)$, for example, a sentence is a sequence of tokens, an image is a sequence of pixels. We fix an order (e.g., left-to-right, row-wise or column-wise, etc).

Factorise $p_{Y|X}(y|x)$ using univariate conditionals.

Examples of factorisation

- 1. full conditional independence $p_{Y|X}(y|x) \stackrel{\text{ind.}}{=} \prod_{n=1}^{N} p_{W|X}(w_n|x)$
- 2. Markov model $p_{Y|X}(y|x) \stackrel{\text{ind.}}{=} \prod_{n=1}^{N} p_{W|XH}(w_n|x, w_{n-k+1:n-1})$
- 3. chain rule $p_{Y|X}(y|x) = \prod_{n=1}^{N} p_{W|XH}(w_n|x, w_{< n})$ aka autoregressive factorisation

Given a flexible-enough family for the conditionals, (3) can identify *any* probability measure, in principle.

See Germain et al. (MADE; 2015) and any decoder-only or encoderdecoder model (Mikolov et al., 2010; Van den Oord et al., 2016; Oord et al., 2016; Vaswani et al., 2017)

Undirected

It's possible to factorise $p_{Y|X}(y|x)$ using unnormalised factors.

For example, a first-order conditional random field $p_{Y|X}(y|x) \propto \prod_{n=1}^{N} \Phi(x, w_{n-1}, w_n)$ uses factors like $\Phi(x, w_{n-1}, w_n) > 0$.

The familiar constraints apply: non-negativity, finite normalisation constant.

Density/mass assessment, sampling may be possible in some cases (eg, first-order CRF) but are intractable in general.

EBMs: with rather flexible NNs, we can parameterise an unnormalised model without an explicit factorisation: $p_{Y|X}(y|x,\theta) \propto NN(x,y;\theta)$.

Honourable mentions

- Sparse continuous distributions (Martins et al., 2022)
- Score matching (implicit generative models) (Vincent, 2011; Song and Ermon, 2019; Song et al., 2020)
- Diffusion processes (Sohl-Dickstein et al., 2015; Kingma et al., 2021)

We will cover score matching and diffusion in this module.

Summary

There are various ways to prescribe distributions both univariate and multivariate.

- Predict parameters for known pdfs and cdfs: we predict some finite (typically small) number of parameters, and evaluate the mass/density of an outcome using a known function.
- For more flexibility we construct novel pdfs or cdfs by predicting unnormalised densities or parameterising flows and simulators.
- For multivariate and structured data we typically exploit a factorisation into simpler distributions (NNs are particularly good at representing complex conditioning contexts).

There are various tradeoffs: is mass/density assessment tractable? can we sample? do we need backward passes? do we need to approximate normalisation constants?

Beyond

There's a lot more to what I said today.

- Big open problems involving parameter estimation, probabilistic inference, and model criticism.
- Creative applications for uncertainty representation (e.g., out-of-domain or error detection, controllable generation).
- Challenging data that may require novel tools.
- Augment our uncertainty representation to include uncertainty about parameters and model family.

If you are excited about LLMs, they too predict a representation of uncertainty: about the response given the prompt.

They exploit an autoregressive factorisation of the conditional distribution, their factors are simple Categorical distributions over a vocabulary of tokens.

Decision making, probabilistic inference, disentanglement learning, representation of epistemic uncertainty, and statistical evaluation are big challenges in that space. See for example our work on decision making (Eikema and Aziz, 2020, 2022) and statistical evaluation (Barkhof and Aziz, 2022; Baan et al., 2022; Giulianelli et al., 2023; Ilia and Aziz, 2024).

References I

Joris Baan, Wilker Aziz, Barbara Plank, and Raquel Fernandez. Stop measuring calibration when humans disagree. In *Proceedings of the* 2022 Conference on Empirical Methods in Natural Language Processing, pages 1892–1915, Abu Dhabi, United Arab Emirates, December 2022. Association for Computational Linguistics. URL https://aclanthology.org/2022.emnlp-main.124.

Joris Baan, Nico Daheim, Evgenia Ilia, Dennis Ulmer, Haau-Sing Li, Raquel Fernández, Barbara Plank, Rico Sennrich, Chrysoula Zerva, and Wilker Aziz. Uncertainty in natural language generation: From theory to applications, 2023.

References II

Joris Baan, Raquel Fernández, Barbara Plank, and Wilker Aziz.
Interpreting predictive probabilities: Model confidence or human label variation? In Yvette Graham and Matthew Purver, editors, *Proceedings of the 18th Conference of the European Chapter of the Association for Computational Linguistics (Volume 2: Short Papers)*, pages 268–277, St. Julian's, Malta, March 2024. Association for Computational Linguistics. URL https://aclanthology.org/2024.eacl-short.24.
Claartje Barkhof and Wilker Aziz. Statistical model criticism of variational

auto-encoders. arXiv preprint arXiv:2204.03030, 2022.

Léon Bottou. Large-scale machine learning with stochastic gradient descent. In *Proceedings of COMPSTAT'2010*, pages 177–186. Springer, 2010.

References III

Nicola De Cao, Wilker Aziz, and Ivan Titov. Block neural autoregressive flow. In Ryan P. Adams and Vibhav Gogate, editors, *UAI*, volume 115 of *Proceedings of machine learning research*, pages 1263–1273, Tel Aviv, Israel, July 2020. PMLR. URL

http://proceedings.mlr.press/v115/de-cao20a.html.tex.pdf: http://proceedings.mlr.press/v115/de-cao20a/de-cao20a.pdf.

Bruno De Finetti. *Theory of probability: a critical introductory treatment*. 1974.

D. Dubois and H. Prade. An Introduction to Possibilistic and Fuzzy Logics, page 742–761. Morgan Kaufmann Publishers Inc., San Francisco, CA, USA, 1990. ISBN 1558601252.

References IV

Bryan Eikema and Wilker Aziz. Is MAP Decoding All You Need? The Inadequacy of the Mode in Neural Machine Translation. arXiv:2005.10283 [cs], May 2020. URL http://arxiv.org/abs/2005.10283. arXiv: 2005.10283.

Bryan Eikema and Wilker Aziz. Sampling-based approximations to minimum Bayes risk decoding for neural machine translation. In Yoav Goldberg, Zornitsa Kozareva, and Yue Zhang, editors, Proceedings of the 2022 Conference on Empirical Methods in Natural Language Processing, pages 10978–10993, Abu Dhabi, United Arab Emirates, December 2022. Association for Computational Linguistics. doi: 10.18653/v1/2022.emnlp-main.754. URL

https://aclanthology.org/2022.emnlp-main.754.

References V

António Farinhas, Wilker Aziz, Vlad Niculae, and Andre Martins. Sparse communication via mixed distributions. In *International Conference on Learning Representations*, 2022. URL https://openreview.net/forum?id=WAid50QschI.

Nir Friedman and Joseph Y. Halpern. Plausibility measures and default reasoning. In *Proceedings of the Thirteenth National Conference on Artificial Intelligence - Volume 2*, AAAI'96, page 1297–1304. AAAI Press, 1996. ISBN 026251091X.

Andrew Gelman, John Carlin, Hal Stern, David Dunson, Aki Vehtari, and Donald Rubin. *Bayesian Data Analysis*. Chapman and Hall/CRC, third edition, 2013.

Mathieu Germain, Karol Gregor, Iain Murray, and Hugo Larochelle. Made: Masked autoencoder for distribution estimation. In *International conference on machine learning*, pages 881–889. PMLR, 2015.

References VI

Mario Giulianelli, Joris Baan, Wilker Aziz, Raquel Fernández, and Barbara Plank. What comes next? evaluating uncertainty in neural text generators against human production variability. In Houda Bouamor, Juan Pino, and Kalika Bali, editors, *Proceedings of the 2023 Conference on Empirical Methods in Natural Language Processing*, pages 14349–14371, Singapore, December 2023. Association for Computational Linguistics. doi: 10.18653/v1/2023.emnlp-main.887. URL https://aclanthology.org/2023.emnlp-main.887.

Moisés Goldszmidt and Judea Pearl. Rank-based systems: A simple approach to belief revision, belief update, and reasoning about evidence and actions. In *Proceedings of the Third International Conference on Principles of Knowledge Representation and Reasoning*, KR'92, page 661–672, San Francisco, CA, USA, 1992. Morgan Kaufmann Publishers Inc. ISBN 1558602623.

References VII

Ian Goodfellow, Jean Pouget-Abadie, Mehdi Mirza, Bing Xu, David Warde-Farley, Sherjil Ozair, Aaron Courville, and Yoshua Bengio. Generative adversarial nets. Advances in neural information processing systems, 27, 2014.

Ian Hacking. The emergence of probability: A philosophical study of early ideas about probability, induction and statistical inference. Cambridge University Press, 1975.

Joseph Y Halpern. Reasoning about uncertainty. MIT press, 2017.

Jaakko Hintikka. *Modality as referential multiplicity*. Filosofisen Yhdistyksen vuosikirja, 1957.

Jaakko Hintikka. Modality and quantification. *Theoria*, 27(3):119–128, 1961.

References VIII

Chin-Wei Huang, David Krueger, Alexandre Lacoste, and Aaron Courville. Neural Autoregressive Flows. In International Conference on Machine Learning, pages 2078–2087. PMLR, July 2018. URL http://proceedings.mlr.press/v80/huang18d.html. ISSN: 2640-3498.

Ferenc Huszár. Variational inference using implicit distributions. *arXiv* preprint arXiv:1702.08235, 2017.

Evgenia Ilia and Wilker Aziz. Predict the next word: <humans exhibit uncertainty in this task and language models _____>. In Yvette Graham and Matthew Purver, editors, Proceedings of the 18th Conference of the European Chapter of the Association for Computational Linguistics (Volume 2: Short Papers), pages 234–255, St. Julian's, Malta, March 2024. Association for Computational Linguistics. URL https://aclanthology.org/2024.eacl-short.22.

References IX

Diederik Kingma, Tim Salimans, Ben Poole, and Jonathan Ho. Variational diffusion models. *Advances in neural information processing systems*, 34:21696–21707, 2021.

Diederik P. Kingma and Max Welling. Auto-Encoding Variational Bayes. In Yoshua Bengio and Yann LeCun, editors, 2nd International Conference on Learning Representations, ICLR 2014, Banff, AB, Canada, April 14-16, 2014, Conference Track Proceedings, 2014. URL http://arxiv.org/abs/1312.6114.

Durk P Kingma, Tim Salimans, Rafal Jozefowicz, Xi Chen, Ilya Sutskever, and Max Welling. Improved Variational Inference with Inverse Autoregressive Flow. In D. D. Lee, M. Sugiyama, U. V. Luxburg, I. Guyon, and R. Garnett, editors, *Advances in Neural Information Processing Systems 29*, pages 4743–4751. Curran Associates, Inc., 2016. URL http://papers.nips.cc/paper/

References X

6581-improved-variational-inference-with-inverse-autoregres pdf.

Daphne Koller and Nir Friedman. *Probabilistic Graphical Models*. MIT Press, 2009.

Andrey N. Kolmogorov. *Foundations of the Theory of Probability*. Chelsea Pub Co, 2 edition, June 1960.

Dennis V Lindley. Understanding uncertainty. John Wiley & Sons, 2013.

Andre Martins and Ramon Astudillo. From Softmax to Sparsemax: A Sparse Model of Attention and Multi-Label Classification. In Maria Florina Balcan and Kilian Q. Weinberger, editors, *Proceedings of The 33rd International Conference on Machine Learning*, volume 48 of *Proceedings of Machine Learning Research*, pages 1614–1623, New York, New York, USA, June 2016. PMLR. URL

http://proceedings.mlr.press/v48/martins16.html.

References XI

André FT Martins, Marcos Treviso, António Farinhas, Pedro MQ Aguiar, Mário AT Figueiredo, Mathieu Blondel, and Vlad Niculae. Sparse continuous distributions and fenchel-young losses. *The Journal of Machine Learning Research*, 23(1):11728–11801, 2022.

Christopher Menzel. Possible Worlds. In Edward N. Zalta and Uri Nodelman, editors, *The Stanford Encyclopedia of Philosophy*.
Metaphysics Research Lab, Stanford University, Summer 2023 edition, 2023.

Tomas Mikolov, Martin Karafiát, Lukas Burget, Jan Cernockỳ, and Sanjeev Khudanpur. Recurrent neural network based language model. In *Interspeech*, volume 2, pages 1045–1048. Makuhari, 2010.

Vlad Niculae and Mathieu Blondel. A regularized framework for sparse and structured neural attention. *Advances in neural information processing systems*, 30, 2017.

References XII

- Vlad Niculae, Caio F Corro, Nikita Nangia, Tsvetomila Mihaylova, and André FT Martins. Discrete latent structure in neural networks. *arXiv* preprint arXiv:2301.07473, 2023.
- Aaron van den Oord, Sander Dieleman, Heiga Zen, Karen Simonyan, Oriol Vinyals, Alex Graves, Nal Kalchbrenner, Andrew Senior, and Koray Kavukcuoglu. Wavenet: A generative model for raw audio. arXiv preprint arXiv:1609.03499, 2016.
- George Papamakarios, Eric Nalisnick, Danilo Jimenez Rezende, Shakir Mohamed, and Balaji Lakshminarayanan. Normalizing Flows for Probabilistic Modeling and Inference. *arXiv:1912.02762 [cs, stat]*, December 2019. URL http://arxiv.org/abs/1912.02762. arXiv: 1912.02762.

References XIII

Ben Peters, Vlad Niculae, and André F. T. Martins. Sparse Sequence-to-Sequence Models. In *Proceedings of the 57th Annual Meeting of the Association for Computational Linguistics*, pages 1504–1519, Florence, Italy, July 2019. Association for Computational Linguistics. doi: 10.18653/v1/P19-1146. URL https://www.aclweb.org/anthology/P19-1146.

Barbara Plank. The "problem" of human label variation: On ground truth in data, modeling and evaluation. In *Proceedings of the 2022 Conference on Empirical Methods in Natural Language Processing*, pages 10671–10682, Abu Dhabi, United Arab Emirates, December 2022. Association for Computational Linguistics. URL

https://aclanthology.org/2022.emnlp-main.731.

Frank Plumpton Ramsey. *The foundations of mathematics and other logical essays*. K. Paul, Trench, Trubner & Company, Limited, 1931.

References XIV

Danilo Rezende and Shakir Mohamed. Variational inference with normalizing flows. In *International conference on machine learning*, pages 1530–1538. PMLR, 2015.

Danilo Jimenez Rezende, Shakir Mohamed, and Daan Wierstra. Stochastic Backpropagation and Approximate Inference in Deep Generative Models. In Eric P. Xing and Tony Jebara, editors, *Proceedings of the 31st International Conference on Machine Learning*, volume 32 of *Proceedings of Machine Learning Research*, pages 1278–1286, Bejing, China, June 2014. PMLR. URL

http://proceedings.mlr.press/v32/rezende14.html. lssue: 2.

Herbert Robbins and Sutton Monro. A stochastic approximation method. *The annals of mathematical statistics*, pages 400–407, 1951. Publisher: JSTOR.

References XV

John Schulman, Nicolas Heess, Theophane Weber, and Pieter Abbeel. Gradient estimation using stochastic computation graphs. In *Advances in neural information processing systems*, pages 3528–3536, 2015.

Glenn Shafer. A Mathematical Theory of Evidence. Princeton University Press, 1976. ISBN 9780691100425. URL http://www.jstor.org/stable/j.ctv10vm1qb.

Jascha Sohl-Dickstein, Eric Weiss, Niru Maheswaranathan, and Surya Ganguli. Deep unsupervised learning using nonequilibrium thermodynamics. In *International conference on machine learning*, pages 2256–2265. PMLR, 2015.

Yang Song and Stefano Ermon. Generative modeling by estimating gradients of the data distribution. *Advances in neural information processing systems*, 32, 2019.

References XVI

- Yang Song, Sahaj Garg, Jiaxin Shi, and Stefano Ermon. Sliced score matching: A scalable approach to density and score estimation. In *Uncertainty in Artificial Intelligence*, pages 574–584. PMLR, 2020.
- Aaron Van den Oord, Nal Kalchbrenner, Lasse Espeholt, Oriol Vinyals, Alex Graves, et al. Conditional image generation with pixelcnn decoders. *Advances in neural information processing systems*, 29, 2016.
- Ashish Vaswani, Noam Shazeer, Niki Parmar, Jakob Uszkoreit, Llion Jones, Aidan N Gomez, Łukasz Kaiser, and Illia Polosukhin. Attention is all you need. *Advances in neural information processing systems*, 30, 2017.
- Pascal Vincent. A connection between score matching and denoising autoencoders. *Neural computation*, 23(7):1661–1674, 2011.